

## Quantitative Genetic Models

### Least Squares Genetic Model

- partition of effects

$$P = G + E + G \times E$$

$P$  is phenotypic effect

$G$  is genetic effect

$E$  is environmental effect

$G \times E$  is interaction effect

- partition of variance components

$$Var(P) = Var(G) + 2Cov(G, E) + Var(E)$$

ignoring  $G \times E$

$$Var(P) = Var(G) + Var(E)$$

further ignoring  $Cov(G, E)$

©ZB Zeng & BS Yandell

4.1

February 8, 2001

## Hardy-Weinberg (1908) Principle

change of allele & genotype frequency  
over generations

- "ideal" assumptions:
  - 1 large (infinite) population
  - 2 random mating
  - (no inbreeding or random drift)
  - 3 no natural or artificial selection
  - 4 no mutation
  - 5 closed to migration
  - 6 discrete generations

- sexual reproduction  $\Rightarrow$  1:1 gamete ratio
- 7 separate sexes

©ZB Zeng & BS Yandell

4.2

February 8, 2001

## Sexual Reproduction to H-W Proportions

genotype	female	male
$B$	$p_{1f}$	$p_{1m}$
$b$	$p_{2f}$	$p_{2m}$

genotype	$BB$	$Bb$	$bb$
frequency	$p_{1f}p_{1m}$	$p_{1f}p_{2m} + p_{2f}p_{1m}$	$p_{2f}p_{2m}$

**autosomal locus: same for male, female**

$$\begin{aligned} p_1 &= p_{BB} + p_{Bb}/2 = p_{1f}p_{1m} + [p_{1f}p_{2m} + p_{2f}p_{1m}]/2 \\ &= [p_{1f}(p_{1m} + p_{2m}) + p_{1m}(p_{1f} + p_{2f})]/2 \\ &= [p_{1f} + p_{1m}]/2 \end{aligned}$$

## next and subsequent generations

genotype	$BB$	$Bb$	$bb$
frequency	$p_1^2$	$2p_1p_2$	$p_2^2$

©ZB Zeng & BS Yandell

4.3

February 8, 2001

## Hardy-Weinberg Equilibrium

- only 1 generation to remove sex differences
- only 2 generations to stabilize proportions
- random mating can stabilize proportions subjected to selection in 1-2 generations

## Hardy-Weinberg Disequilibrium

	$B$	$b$	
$B$	$p_B^2 + D_B$	$p_Bp_b - D_B$	$p_B = p_1$
$b$	$p_bp_B - D_B$	$p_b^2 + D_B$	$p_b = p_2$
	$p_B = p_1$	$p_b = p_1$	1

©ZB Zeng & BS Yandell

4.4

February 8, 2001

### Test of Hardy-Weinberg Disequilibrium

- chi-square test

$$X^2 = \sum (\text{obs-exp})^2/\text{exp}$$

$$X^2 = D_B^2 \left[ \frac{N_{BB}}{p_B^2} + \frac{2N_{Bb}}{p_B p_b} + \frac{N_{bb}}{p_b^2} \right]$$

- log likelihood test

$$G = -2 \sum \text{obs} \times \log(\text{exp}/\text{obs})$$

$$\begin{aligned} G = & 2[N_{BB} \log \left( 1 + \frac{D_B^2}{p_B^2} \right) + N_{Bb} \log \left( \frac{D_B^2}{p_B p_b} \right) \\ & + N_{bb} \log \left( 1 - \frac{D_B^2}{p_b^2} \right)] \end{aligned}$$

©ZB Zeng & BS Yandell

4.5

February 8, 2001

### Hardy-Weinberg Equilibrium

gametes	<i>B</i>	<i>b</i>	total
<i>B</i>	$p^2$	$p(1-p)$	$p_1 = p$
<i>b</i>	$(1-p)p$	$(1-p)^2$	$p_2 = 1-p$
total	$p_1 = p$	$p_2 = 1-p$	1

genotype frequency	<i>BB</i>	<i>Bb</i>	<i>bb</i>
	$f_{BB} = (p_1)^2$	$f_{Bb} = 2p_1 p_2$	$f_{bb} = (p_2)^2$

### Genetic Variance Components for a Locus

Genotype	value	common	Fisher
<i>BB</i>	$G_{11}$	$m + a$	$\mu + 2a_1 + d_{11}$
<i>Bb</i>	$G_{12} = G_{21}$	$m + d$	$\mu + a_1 + a_2 + d_{12}$
<i>bb</i>	$G_{22}$	$m - a$	$\mu + 2a_2 + d_{22}$

$$\mu = \sum_g f_g G_g = \sum_{i,j} p_i p_j G_{ij} = m + (p_1 - p_2)a + 2p_1 p_2 d$$

©ZB Zeng & BS Yandell

4.6

©ZB Zeng & BS Yandell

4.6

February 8, 2001

### Additive and Dominance Variance Components Mather-Jinks Partition

$$G_{11} = m + a, G_{12} = m + d, G_{22} = m - a$$

$$\begin{aligned} \sigma_G^2 &= \sum_g f_g (G_g - \mu)^2 = \sum_{i,j} p_i p_j G_{ij}^2 - \mu^2 \\ &= p_1^2 a^2 + 2p_1 p_2 d^2 + p_2^2 (-a)^2 - (\mu - m)^2 \\ &= 2p_1 p_2 [a + (p_2 - p_1)d]^2 + 4p_1^2 p_2^2 d^2 \\ &= \sigma_a^2 + \sigma_d^2 \end{aligned}$$

#### additive variance component

$$\sigma_a^2 = 2p_1 p_2 [a + (p_2 - p_1)d]^2$$

$$\text{dominant variance component } \sigma_d^2 = 4p_1^2 p_2^2 d^2$$

©ZB Zeng & BS Yandell

4.7

February 8, 2001

### Fisher's Least Square Partition of Genetic Variance Components

$$\begin{aligned} G_{ij} &= \mu + a_i + a_j + d_{ij} \\ &= \mu + a_1 N_1(ij) + a_2 N_2(ij) + d_{ij} \end{aligned}$$

$N_1, N_2$  = number copies of allele (1=B, 2=b)

#### parameters

$$\mu = \bar{G}_{..} = \sum_{i,j} p_i p_j G_{ij}$$

$$a_i = \sum_j p_j G_{ij} - \mu = \bar{G}_{i.} - \bar{G}_{..}$$

$$a_j = \bar{G}_{.j} - \bar{G}_{..}$$

$$d_{ij} = G_{ij} - a_i - a_j - \mu = G_{ij} - \bar{G}_{i.} - \bar{G}_{.j} + \bar{G}_{..}$$

which satisfy the conditions

$$\sum_i p_i a_i = \sum_{i,j} p_i p_j d_{ij} = \sum_i p_i d_{ij} = \sum_j p_j d_{ij} = 0$$

©ZB Zeng & BS Yandell

4.8

February 8, 2001

### LS Reference Mean and Additive Effects

$$\begin{aligned}\mu &= p_1^2 G_{11} + 2p_1 p_2 G_{12} + p_2^2 G_{22} \\ a_1 &= p_2[p_1(G_{11} - G_{12}) + p_2(G_{12} - G_{22})] \\ a_2 &= -p_1[p_1(G_{11} - G_{12}) + p_2(G_{12} - G_{22})]\end{aligned}$$

### relation to Mather-Jinks

$$\begin{aligned}\mu &= m + a(p_1^2 - p_2^2) + 2dp_1p_2 \\ a_1 &= p_2[a + d(p_2 - p_1)] \\ a_2 &= -p_1[a + d(p_2 - p_1)] \\ a_1 - a_2 &= a + d(p_2 - p_1)\end{aligned}$$

### LS Dominance Effects

$$\begin{aligned}d_{11} &= G_{11} - \mu - 2a_1 = -p_2^2[2G_{12} - G_{11} - G_{22}] \\ d_{12} &= d_{21} = G_{12} - \mu - a_1 - a_2 \\ &= p_1p_2[2G_{12} - G_{11} - G_{22}] \\ d_{22} &= G_{22} - \mu - 2a_2 = -p_1^2[2G_{12} - G_{11} - G_{22}]\end{aligned}$$

©ZB Zeng & BS Yandell

4.9

February 8, 2001

### LS genetic variance components

$$\begin{aligned}\sigma_G^2 &= \sum_{i,j} p_i p_j (G_{ij} - \mu)^2 = \sum_{i,j} p_i p_j (a_i + a_j + d_{ij})^2 \\ &= 2 \sum_i p_i a_i^2 + \sum_i p_i p_j d_{ij}^2 \\ &= \sigma_a^2 + \sigma_d^2\end{aligned}$$

### Additive Variance

$$\begin{aligned}\sigma_a^2 &= 2 \sum_i p_i a_i^2 = 2[p_1 a_1^2 + p_2 a_2^2] \\ &= 2p_1 p_2 [p_1(G_{11} - G_{12}) + p_2(G_{12} - G_{22})]^2 \\ &= 2p_1 p_2 [p_1(a - d) + p_2(a + d)]^2 \\ &= 2p_1 p_2 [a + (p_2 - p_1)d]^2\end{aligned}$$

### Dominance Variance

$$\begin{aligned}\sigma_d^2 &= \sum_{i,j} p_i p_j d_{ij}^2 = p_1^2 d_{11}^2 + 2p_1 p_2 d_{12}^2 + p_2^2 d_{22}^2 \\ &= p_1^2 p_2^2 [2G_{12} - G_{11} - G_{22}]^2 \\ &= 4p_1^2 p_2^2 d^2\end{aligned}$$

©ZB Zeng & BS Yandell

4.10

February 8, 2001

### Hardy-Weinberg Disequilibrium

genotype	$BB$	$Bb$	$bb$
frequency	$P_{BB}$	$P_{Bb}$	$P_{bb}$
H-W equilibrium	$p_B^2$	$2p_B p_b$	$p_b^2$

with  $p_B = P_{BB} + \frac{1}{2}P_{Bb}$  and  $p_b = P_{bb} + \frac{1}{2}P_{Bb}$

### disequilibrium coefficients

$$\begin{aligned}P_{BB} &= p_B^2 + D_{BB} \\ P_{Bb} &= 2p_B p_b - 2D_{Bb} \\ P_{bb} &= p_b^2 + D_{bb}\end{aligned}$$

constraints of gene and genotype frequencies

$$\begin{aligned}P_{BB} + P_{Bb} + P_{bb} &= 1 \Rightarrow \sum D_{ij} = 0 \\ p_{BB} + p_{Bb}/2 &= p_B \Rightarrow D_{BB} = D_{Bb} = D_{bb}\end{aligned}$$

$D_B$  measures H-W disequilibrium for locus

$$\begin{aligned}P_{BB} &= p_B^2 + D_B \\ P_{Bb} &= 2p_B p_b - 2D_B \\ P_{bb} &= p_b^2 + D_B\end{aligned}$$

©ZB Zeng & BS Yandell

4.11

February 8, 2001

### Linkage disequilibrium

- Let us consider the distribution of gametes for a two-allele and two-locus model:

Gamete	$AB$	$Ab$	$aB$	$ab$
Frequency	$P_{AB}$	$P_{Ab}$	$P_{aB}$	$P_{ab}$
L.-E.	$p_A p_B$	$p_A p_b$	$p_a p_B$	$p_a p_b$

Define disequilibrium coefficients as:

$$\begin{aligned}P_{AB} &= p_A p_B + D_{AB} \\ P_{Ab} &= p_A p_b + D_{Ab} \\ P_{aB} &= p_a p_B + D_{aB} \\ P_{ab} &= p_a p_b + D_{ab}\end{aligned}$$

constraints of gene and genotype frequencies

$$D_{AB} + D_{Ab} + D_{aB} + D_{ab} = 0$$

$$p_A = P_{AB} + P_{Ab} = p_A + D_{AB} + D_{Ab}$$

$$\Rightarrow D_{AB} = -D_{Ab} \text{ and } D_{ab} = -D_{aB},$$

©ZB Zeng & BS Yandell

4.12

February 8, 2001

In general,

$$\sum_i D_{ij} = \sum_j D_{ij} = 0.$$

Thus

$$\begin{aligned} P_{AB} &= p_A p_B + D_{AB} \\ P_{Ab} &= p_A p_b - D_{AB} \\ P_{aB} &= p_a p_B - D_{AB} \\ P_{ab} &= p_a p_b + D_{AB} \end{aligned}$$

- Correlation coefficient between two loci:

$$\gamma = \frac{D_{AB}}{\sqrt{p_A p_B p_a p_b}}$$

- Relationship between linkage disequilibrium and recombination frequency: For both backcross and F<sub>2</sub> populations, segregating gametes are produced from F<sub>1</sub>

©ZB Zeng & BS Yandell

4.13

February 8, 2001

- The linkage disequilibrium in these populations are determined by the gametes of F<sub>1</sub>. F<sub>1</sub> produces four gametes in the following proportion:

Gamete	AB	Ab	aB	ab
Frequency	$\frac{1-r}{2}$	$\frac{r}{2}$	$\frac{r}{2}$	$\frac{1-r}{2}$

$r$  = recombination frequency ( $0 < r < 0.5$ ).

$$\begin{aligned} p_A &= p_B = p_a = p_b = \frac{1}{2} \\ D_{AB} &= P_{AB} - p_A p_B = \frac{1-r}{2} - \frac{1}{4} = \frac{1}{4}(1-2r) \\ \text{thus} \end{aligned}$$

$$\gamma = \frac{D_{AB}}{\sqrt{p_A p_B p_a p_b}} = \frac{(1-2r)/4}{\sqrt{1/16}} = 1-2r$$

©ZB Zeng & BS Yandell

4.14

February 8, 2001

4.14

February 8, 2001

### Linkage disequilibrium

B	b	
A	$p_A p_B + D_{AB}$	$p_A p_b - D_{AB}$
a	$p_a p_B - D_{AB}$	$p_a p_b + D_{AB}$
$p_B$	$p_b = 1 - p_B$	1

### correlation between loci

$$X = 1(A), 0(a); Y = 1(B), 0(b)$$

$$E(X) = p_A \times 1 + p_a \times 0 = p_A$$

$$\begin{aligned} Var(X) &= p_A \times (1-p_A)^2 + p_a \times (0-p_A)^2 \\ &= p_A p_a \end{aligned}$$

$$\begin{aligned} Cov(X, Y) &= \sum_{x,y} P_{xy} \times (x - E(X))(y - E(Y)) \\ &= (p_A p_B + D_{AB}) \times (1-p_A)(1-p_B) + \\ &\quad (p_A p_b - D_{AB}) \times (1-p_A)(0-p_B) + \dots \\ &= D_{AB} \end{aligned}$$

$$\begin{aligned} \gamma &= Corr(X, Y) = \frac{Cov(X, Y)}{\sqrt{Var(X)Var(Y)}} \\ &= \frac{D_{AB}}{\sqrt{p_A p_B p_a p_b}} \end{aligned}$$

©ZB Zeng & BS Yandell

4.15

February 8, 2001

### Linkage disequilibrium and partition of genetic variance

Consider two loci, no epistasis  
locus A: i, j; locus B: k, l

### least square genetic model

$$G_{jl}^{ik} = \mu + a^i + a^k + a_j + a_l + d_j^i + d_l^k$$

ik = gamete from parent 1 (superscript)

jl = gamete from parent 2 (subscript)

### genotype frequency

$P_{jl}^{ik}$  = frequency of  $G_{jl}^{ik}$  in population

Hardy-Weinberg equilibrium:  $P_{jl}^{ik} = P_{ik} P_{jl}$

linkage disequilibrium:  $P_{ik} = p_i p_k + D_{ik}$   
constraints

$$\sum_i p_i a_i = \sum_i p_i d_j^i = \sum_j p_j d_j^i = \sum_i D_{ik} = \sum_k D_{ik} = 0$$

©ZB Zeng & BS Yandell

4.16

February 8, 2001

**mean of the population  
effects**

$$G_{jl}^{ik} = \mu + a^i + a^k + a_j + a_l + d_j^i + d_l^k$$

constraints

$$\sum_i p_i a_i = \sum_i p_i d_j^i = \sum_j p_j d_j^i = \sum_i D_{ik} = \sum_k D_{ik} = 0$$

$$\begin{aligned} \bar{G}_{..} &= \sum_{i,k} \sum_{j,l} P_{jl}^{ik} G_{jl}^{ik} = \sum_{i,k} \sum_{j,l} P_{ik} P_{jl} G_{jl}^{ik} \\ &= \sum_{i,k} \sum_{j,l} (p_i p_k + D_{ik}) (p_j p_l + D_{jl}) G_{jl}^{ik} \\ &= \sum_{i,k} \sum_{j,l} p_i p_k p_j p_l G_{jl}^{ik} = \mu = \bar{G}_{..} \end{aligned}$$

linkage disequilibrium does not affect mean  
(not true with epistasis)

©ZB Zeng & BS Yandell

4.17

February 8, 2001

**variance of the population**

$$\begin{aligned} \sigma_G^2 &= \sum_{i,k} \sum_{j,l} P_{ik} P_{jl} (G_{jl}^{ik} - \mu)^2 \\ &= \sum_{i,k} \sum_{j,l} (p_i p_k + D_{ik}) (p_j p_l + D_{jl}) (G_{jl}^{ik} - \mu)^2 \\ &= \sum_i p_i (a^i)^2 + \sum_k p_k (a^k)^2 + \sum_j p_j a_j^2 + \sum_l p_l a_l^2 \\ &\quad + 2 \sum_{i,k} D_{ik} a^i a^k + 2 \sum_{j,l} D_{jl} a_j a_l \\ &\quad + \sum_{i,j} p_i p_j d_j^i d_l^k + \sum_{k,l} p_k p_l d_l^k d_l^k + 2 \sum_{i,k} \sum_{j,l} D_{ik} D_{jl} d_j^i d_l^k \end{aligned}$$

**linkage disequilibrium**

variance contains covariances

between additive effects  $a^i$  and  $a^k$ ,  $a_j$  and  $a_l$   
between dominance effects  $d_j^i$  and  $d_l^k$

**H-W equilibrium**

no covariance between  $a^i$  and  $d_j^i$   
inbreeding yields some covariance

©ZB Zeng & BS Yandell

4.18

February 8, 2001

**Genetic model for BC and F<sub>2</sub>**

genotype effects at loci  $B_i, i = 1, 2, \dots, m$

genotype	$B_i B_i$	$B_i b_i$	$b_i b_i$
effect	$a_i$	$d_i$	$-a_i$

$P_1$  population fixed with alleles  $B_i$

$P_2$  population fixed with allele  $b_i$

ignore epistasis

- examine trait distribution
- familiarize with various crosses
- underlying model for QTL
- get ready for epistasis

©ZB Zeng & BS Yandell

4.19

February 8, 2001

**phenotypic value of individual  $k$**

$P_1$  population:

$$y_k = \mu + \sum_{i=1}^m a_i + e_k, \quad e_k \sim N(0, \sigma_e^2)$$

mean and variance of trait in  $P_1$ :

$$\mu_{P_1} = \mu + \sum_{i=1}^m a_i, \quad \sigma_{P_1}^2 = \sigma_e^2$$

Similarly for  $P_2$ :

$$y_k = \mu - \sum_{i=1}^m a_i + e_k, \quad \mu_{P_2} = \mu - \sum_{i=1}^m a_i, \quad \sigma_{P_2}^2 = \sigma_e^2$$

$F_1 = P_1 \times P_2$  population:

$$y_k = \mu + \sum_{i=1}^m d_i + e_k, \quad \mu_{F_1} = \mu + \sum_{i=1}^m d_i, \quad \sigma_{F_1}^2 = \sigma_e^2$$

©ZB Zeng & BS Yandell

4.20

February 8, 2001

**Backcross**  $B_1 = P_1 \times F_1$

$$\begin{aligned} y_k &= \mu + \sum_{i=1}^m [x_{ik}a_i + (1 - x_{ik})d_i] + e_k \\ &= \mu + \sum_{i=1}^m d_i + \sum_{i=1}^m x_{ik}(a_i - d_i) + e_k \\ x_{ik} &= \begin{cases} 1 & \text{if a } B_i \text{ allele in } k\text{'s } F_1 \text{ gamete} \\ 0 & \text{if a } b_i \text{ allele} \end{cases} \\ \mathcal{E}(x_{ik}) &= 1/2, \quad \text{Var}(x_{ik}) = 1/4 \quad \text{Cov}(x_{ik}, x_{jk}) = (1 - 2r_{ij})/4 \end{aligned}$$

mean and variance

$$\begin{aligned} \mu_{B_1} &= \mu + \sum_{i=1}^m (a_i + d_i)/2 \\ \sigma_{B_1}^2 &= \sum_{i=1}^m \frac{(a_i - d_i)^2}{4} + \sum_{i \neq j} \frac{(1 - 2r_{ij})(a_i - d_i)(a_j - d_j)}{4} + \sigma_e^2 \end{aligned}$$

similarly for  $B_2 = P_2 \times F_1$

**$F_2 = F_1 \times F_1$  Population**

$$\begin{aligned} y_k &= \mu + \sum_{i=1}^m [x_{ik1}x_{ik2}a_i + (1 - x_{ik1})(1 - x_{ik2})(-a_i) \\ &\quad + [x_{ik1}(1 - x_{ik2}) + (1 - x_{ik1})x_{ik2}]d_i] + e_k \\ &= \mu - \sum_{i=1}^m a_i + \sum_{i=1}^m [x_{ik}.(a_i + d_i) - 2x_{ik1}x_{ik2}d_i] + e_k \\ x_{ik1} &= \begin{cases} 1 & \text{if } B_i \text{ allele in first gamete from } F_1 \\ 0 & \text{if } b_i \text{ allele} \end{cases} \end{aligned}$$

$$\mathcal{E}(x_{ik1}) = \frac{1}{2}, \quad \text{Var}(x_{ik1}) = \frac{1}{4}$$

$$\text{Cov}(x_{ik1}, x_{jk1}) = \frac{1}{4}(1 - 2r_{ij})$$

and similarly for  $x_{ik2}$

Under the assumption of Hardy-Weinberg equilibrium,  $x_{ik1}$  and  $x_{ik2}$  are independent. With

these conditions, the mean and variance of individuals in  $F_2$  population become

$$\begin{aligned} \mu_{F_2} &= \mu + \frac{1}{2} \sum_{i=1}^m d_i \\ \sigma_{F_2}^2 &= \frac{1}{2} \sum_{i=1}^m a_i^2 + \frac{1}{4} \sum_{i=1}^m d_i^2 + \frac{1}{2} \sum_{i \neq j} (1 - 2r_{ij})a_i a_j \\ &\quad + \frac{1}{4} \sum_{i \neq j} (1 - 2r_{ij})^2 d_i d_j + \sigma_e^2. \end{aligned}$$

first term: additive variance within loci

second: dominance variance within loci

third: additive covariance between loci

fourth: dominance covariance between loci

**Epistasis**

section 4.6 of Zeng's notes  
postponed to chapter 8

**Multiple Interval Mapping**