

CONDITIONAL DISTRIBUTIONS AND MOMENTS

EXAMPLE

Let X and Y have the joint probability mass function specified in the following table.

Joint Probability Mass Function
of X and Y

		x			
		0	1	2	
y	0	0	.2	.1	.3
	1	.2	.4	0	.6
	2	.1	0	0	.1
		.3	.6	.1	1.0

The conditional distribution of Y given $X = 1$ is obtained by restricting attention to the column of values $f_{X,Y}(1, y)$ corresponding to $X = 1$. This needs to be rescaled to distribute probability one.

Conditional Probability Mass Function of Y given $X = 1$

y	0	1	2
$f_{Y X=1}(y)$	$\frac{1}{3}$	$\frac{2}{3}$	0

The **conditional probability mass function** of Y given $X = x$ is defined as

$$f_{Y|X=x}(y) = \frac{f_{X,Y}(x, y)}{f_X(x)}$$

provided $f_X(x) > 0$.

This distribution has moments(conditional) just as any other distribution. For any function $g(\cdot)$, the conditional expectation of $g(Y)$ given $X = x$ is defined as

$$E[Y|X = x] = \sum_y g(y) f_{Y|X=x}(y) \quad \text{discrete case}$$

$$\int_{-\infty}^{\infty} g(y) f_{Y|X=x}(y) \quad \text{continuous case}$$

In particular, with $g(y) = y$,

$$E(Y|X = 1) = 0 \times \frac{1}{3} + 1 \times \frac{2}{3} + 2 \times 0 = \frac{2}{3}$$

In general, the conditional variance is defined as

$$\text{var}(Y|X = x) = \sum_y (y - E[Y|X = x])^2 f_{Y|X=x}(y)$$

$$\begin{aligned}
&= \sum_y (y^2 - 2yE[Y|X = x] + (E[Y|X = x])^2) f_{Y|X=x}(y) \\
&= \sum_y y^2 f_{Y|X=x}(y) - 2 \sum_y y E[Y|X = x] f_{Y|X=x}(y) + (E[Y|X = x])^2 \sum_y f_{Y|X=x}(y) \\
&= \sum_y y^2 f_{Y|X=x}(y) - (E[Y|X = x])^2 = E[Y^2|X = x] - (E[Y|X = x])^2
\end{aligned}$$

Since, in our example,

$$E(Y^2|X = 1) = 0^2 \times \frac{1}{3} + 1^2 \times \frac{2}{3} \times .6 + 2^2 \times 0 = \frac{2}{3}$$

$$\text{we have } \text{var}(Y|X = 1) = \frac{2}{3} - \left(\frac{2}{3}\right)^2 = \frac{2}{9}$$

More Properties of Conditional Expectation

Above, we have shown that the conditional variance

$$\text{var}(Y|X = x) = \sum_y (y - E[Y|X = x])^2 f_{Y|X=x}(y) = E[Y^2|X = x] - (E[Y|X = x])^2$$

Notice that the conditional expectation of Y given $x = x$,

$$E[Y|X = x] = \int_{-\infty}^{\infty} y \frac{f_{X,Y}(x, y)}{f_X(x)} dy$$

is a function of the observed value x of X . Let us now consider X to be random and write $E[Y|X]$ for the random quantity that takes value $E[Y|X = x]$ when $X = x$. We establish two important properties regarding conditional expected values.

Property 1. $E[E[Y|X]] = E[Y]$

Property 2. $E[E[XY|X]] = E[XY]$

First, $E[Y|X]$ has expected value

$$E[E[Y|X]] = \int_{-\infty}^{\infty} E[Y|X = x] f_X(x) dx = \int_{-\infty}^{\infty} \left(\int_{-\infty}^{\infty} y \frac{f_{X,Y}(x, y)}{f_X(x)} dy \right) f_X(x) dx$$

Canceling $f_X(x)$, over the range where it is positive, this last double integral becomes

$$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} y f_{X,Y}(x, y) dx dy = \int_{-\infty}^{\infty} y f_Y(y) dy = E[Y]$$

(When $E[|Y|] < \infty$, the integration can be done in any order).

To establish the second property, we notice that

$$E[XY|X = x] = \int_{-\infty}^{\infty} xy \frac{f_{X,Y}(x, y)}{f_X(x)} dy = x E[Y|X = x]$$

Considering X as random, we write $E[E[XY|X]]$ which has expected value

$$\begin{aligned} E[E[XY|X]] &= \int_{-\infty}^{\infty} E[XY|X=x]f_X(x)dx = \int_{-\infty}^{\infty} \left(\int_{-\infty}^{\infty} xy \frac{f_{X,Y}(x,y)}{f_X(x)} dy \right) f_X(x) dx \\ &= \int_{-\infty}^{\infty} \left(\int_{-\infty}^{\infty} xy f_{X,Y}(x,y) dx dy \right) = E[XY] \end{aligned}$$

as claimed.

EXAMPLE Suppose the joint distribution of X and Y is such that

$$E[X] = 5 \quad E[X^2] = 30 \quad E[Y|X] = 2 + 3X \quad \text{Var}(Y) = 81$$

Find (a) $\text{Cov}(X, Y)$ and (b) $\text{Corr}(X, Y)$.

The answer requires the two properties above.

Solution to Example Using the properties above

$$E[Y] = E[E[Y|X]] = E[2 + 3X] = 2 + 3E[X] = 2 + 3(5) = 17$$

$$\begin{aligned} E[XY] &= E[E[XY|X]] = E[2X + 3X^2] \\ &= 2E[X] + 3E[X^2] = 10 + 3(30) = 100 \end{aligned}$$

Consequently, $\text{Cov}(X, Y) = E[XY] - E[X]E[Y] = 100 - (17)(5) = 15$ and $\text{Var}(X) = E[X^2] - (E[X])^2 = 30 - 5^2 = 5$ so

$$\text{Corr}(X, Y) = \frac{\text{Cov}(X, Y)}{\sqrt{\text{Var}(X)}\sqrt{\text{Var}(Y)}} = \frac{15}{\sqrt{5}\sqrt{81}}$$

The next decomposition of variance is useful in many statistical applications.

Property 3. $\text{var}(Y) = E[\text{var}(Y|X)] + \text{var}(E[Y|X])$

Proof:

Write $Y - \mu_Y = Y - E[Y|X] + E[Y|X] - \mu_Y$ and square

$$(Y - \mu_Y)^2 = (Y - E[Y|X])^2 + (E[Y|X] - \mu_Y)^2 + 2(Y - E[Y|X])(E[Y|X] - \mu_Y)$$

Take expected values with respect to the joint distribution of (X, Y) . The left hand side gives $\text{var}(Y)$. For the first term on the right hand side, we take the iterated expected value by conditioning on X . and using

$$E[(Y - E[Y|X])^2 | X = x] = \text{var}(Y|X = x)$$

Next, taking the expected value over X produces the first term in Property 3. By Property 1 $E[E[Y|X]] = E[Y] = \mu_Y$ so the second term is the conditional variance of $E[Y|X]$. Finally, the conditional expected value of the cross-term

$$\begin{aligned} E [[2(Y - E[Y|X])(E[Y|X] - \mu_Y)|X]] \\ = 2(E[Y|X] - \mu_Y)E [(Y - E[Y|X])|X] = 0 \end{aligned}$$

so the unconditional expected value is also 0 and Property 3 follows.